



## MOMENT INEQUALITIES FOR SOME BIVARIATE AGEING CLASSES

Syed Tahir Hussainy, A. Sathiyaraj and U. Rizwan

Department of Mathematics, Islamiah College, Vaniyambadi 635 752, India.

### A B S T R A C T

### RESEARCH ARTICLE

In this paper, we derive the Moments Inequalities for some bivariate ageing classes

**AMS Subject Classification:** 60K10

#### Keywords:

Moment Inequalities for Bivariate Ageing classes, BIFR, BNBUE, BHNBU, BNUFR, BNBRU, BRNBRU, BRNBRUE, BNBUL, BEBU, BEBUC(2), BEBUCA.

Copyright©2018 *Syed Tahir Hussainy, A. Sathiyaraj and U. Rizwan*. This is an open access article distributed under the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

## 1. Introduction

In this paper we establish inequalities for the moments of bivariate ageing classes, namely, Bivariate Increasing Failure Rate (BIFR), Bivariate New Better than Used in Expectation (BNBUE), Bivariate Harmonic New Better than Used in Expectation (BHNBU), Bivariate New Better than Used in Failure Rate (BNUFR), Bivariate New Better than Renewal Used (BNBRU), Bivariate New Renewal better than Used (BNRBU), Bivariate Renewal New is Better than Renewal Used in Expectation (BRNBRUE), Bivariate New Better than Used in Laplace transform order (BNBUL), Bivariate Exponential Better than Used (BEPU), Bivariate Exponential Better than Used in Convex order (BEBUC(2)), Bivariate Exponential Better than Used in Convex Average (BEBUCA), Certain class of life distributions and their variations have been introduced. The application of these bivariate classes of life distributions can be seen in engineering, social and biological sciences. Reliability analysis have shown a growing interest in modeling, survival data using classification of life distributions based on some aspects of aging.

## 2. Preliminaries

In this section, we give below the definition of various bivariate stochastic ageing classes that are required for further discussion.

**Definition 2.1** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Increasing Failure Rate (BIFR) if

$$\frac{\bar{F}(x+t, y+s)}{\bar{F}(t, s)}$$

is decreasing in  $t$ , whenever  $x, y \geq 0$ .

**Definition 2.2** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate New Better

than Used in Expectation (BNBUE) if

$$\int_0^\infty \int_0^\infty \bar{F}(x+t, y+s) dt ds \leq \bar{F}(x, y) \int_0^\infty \int_0^\infty \bar{F}(t, s) dt ds \text{ for } x, y \geq 0.$$

**Definition 2.3** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Harmonic New Better than Used in Expectation (BHNBU) if

$$\int_t^\infty \int_s^\infty \bar{F}(x, y) dy dx \leq \sim \exp\left[-\frac{t+s}{\sim}\right], \text{ for } t, s \geq 0$$

and  $\sim$  denotes the finite mean given by

$$\sim = \int_0^\infty \int_0^\infty \bar{F}(x, y) dy dx$$

**Definition 2.4** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  having failure rate  $r(x, y)$  is said to have Bivariate New Better than Used in Failure Rate (BNUFR) if

$$r(0,0) \leq r(x, y) \text{ for } x, y \geq 0.$$

**Definition 2.5** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate New Better than Renewal Used (BNBRU) if

$$\int_{y+s}^\infty \int_{x+t}^\infty \bar{F}(u, v) dv du \leq \bar{F}(x, y) \int_t^\infty \int_s^\infty \bar{F}(u, v) dv du \text{ for } x, y, t, s \geq 0.$$

Put

$$\bar{W}(x, y) = \frac{1}{\sim} \int_y^\infty \int_x^\infty \bar{F}(u, v) dv du$$

then the above inequality becomes

$$\bar{W}(x+t, y+s) \leq \bar{F}(x, y) \bar{W}(t, s), \text{ for all } x, y, t, s \geq 0$$

**Definition 2.6** A bivariate random variable  $(X, Y)$  or its

distribution  $\bar{F}(x, y)$  is said to have Bivariate New Renewal Better than Renewal Used (BNRBU) if

$$\bar{F}(x+t, y+s) \leq \bar{F}(x, y)\bar{W}(t, s),$$

where

$$\bar{W}(t, s) = \int_s^\infty \int_t^\infty \bar{F}(u, v) dv du.$$

**Definition 2.7** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Renewal New Better than Used (BRNBU) if

$$\bar{F}_{t,s}(x, y) \leq \bar{W}_F(x, y)$$

where,

$$\bar{F}_{t,s}(x, y) = \frac{\bar{F}(t+x, s+y)}{\bar{F}(t, s)} : \bar{F} > 0$$

$$\bar{W}_F(x, y) = \frac{1}{\sim} \int_x^\infty \int_y^\infty \bar{F}(u, v) dv du \text{ for } x, y \geq 0$$

Here  $\sim$  denotes the mean of the bivariate life distribution  $F$  and is assumed to be finite.

**Definition 2.8** A bivariate random variable  $(x, y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Renewal New is Better than Renewal Used (BRNBRU) if

$$\bar{W}_F(x+t, y+s) \leq \bar{W}_F(x, y)\bar{W}_F(t, s).$$

For all  $x, y, t, s \geq 0$  That is

$$\begin{aligned} \sim \int_{x+s}^\infty \int_{y+s}^\infty \bar{F}(u, v) dv du &\leq \left( \int_x^\infty \int_y^\infty \bar{F}(u, v) dv du \right) \\ &\quad \left( \int_t^\infty \int_s^\infty \bar{F}(u, v) dv du \right), \end{aligned}$$

for all  $x, y, t, s \geq 0$ .

**Definition 2.9** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate New Better than Used in Laplace transform order (BNBUL) if

$$\begin{aligned} \int_0^\infty \int_0^\infty \exp[-\lambda](x+y) \bar{F}(x+t, y+s) dy dx \\ \leq \bar{F}(t, s) \int_0^\infty \int_0^\infty \exp[-\lambda](x+y) \bar{F}(x, y) dy dx, \end{aligned}$$

for all  $x, y > 0$  and  $\lambda \geq 0$ .

**Definition 2.10** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Renewal New is Better than Renewal Used in Expectation (BRNBRUE) if

$$2\sim \int_x^\infty \int_y^\infty \int_u^\infty \int_v^\infty \bar{F}(t, s) ds dt \leq \sim \int_x^\infty \int_y^\infty \bar{F}(t, s) ds dt.$$

**Definition 2.11** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Exponential Better than (BEBU) if

$$\bar{F}(x+t, y+s) \leq \bar{F}(t, s) e^{\frac{(x+y)}{\sim}} \text{ for all } x, t, y, s \geq 0.$$

**Definition 2.12** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Exponential Better than Used in Convex order Two (BEBUC(2)) if

$$\begin{aligned} &\int_u^\infty \int_v^\infty \bar{F}(x+t, y+s) ds dt \\ &\leq \sim \exp \left[ -\left( \frac{x+y}{\sim} \right) \right] \int_u^\infty \int_v^\infty \bar{F}(t, s) ds dt, \end{aligned}$$

for all  $x, y, u, v \geq 0$ .

**Definition 2.13** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Exponential Better than Used Convex order Average (BEBUCA) if

$$\int_0^\infty \int_0^\infty \int_{y+s}^\infty \int_{x+t}^\infty \bar{F}(u, v) dv du dx dy \leq \sim^2 \bar{F}(t, s)$$

Put

$$\bar{W}(x+t, y+s) = \int_{y+s}^\infty \int_{x+t}^\infty \bar{F}(u, v) dv du$$

then the above inequality becomes

$$\int_0^\infty \int_0^\infty \bar{W}(x+t, y+s) dx dy \leq \sim^2 \bar{F}(t, s), \text{ for all } t, s \geq 0.$$

**Definition 2.14** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate Overall Decreasing Life class (BODL) if

$$\int_s^\infty \int_t^\infty \bar{W}(x, y) dx dy \leq \sim \bar{W}(t, s),$$

where

$$\bar{W}(t, s) = \frac{1}{\sim} \int_s^\infty \int_t^\infty \bar{F}(u, v) dv du$$

and

$$\sim = \int_0^\infty \int_0^\infty \bar{F}(u, v) dv du$$

is assumed to be finite.

**Definition 2.15** A bivariate random variable  $(X, Y)$  or its distribution  $\bar{F}(x, y)$  is said to have Bivariate  $r^{th}$  Moment

$$\sim_r = E(X^n Y^n) = n^2 \int_0^\infty \int_0^\infty (xy)^{n-1} \bar{F}(x, y) dy dx.$$

### 3. Moment Inequality

In this section we established some theorems on Moment Inequality. We now present a theorem which is useful for further discussion

**Theorem 3.1** If (i)  $F$  is BIFR with mean  $\sim$  and

$$\bar{G}(x, y) = e^{-\frac{\sqrt{x^2 + y^2}}{\sim}}$$

(ii)  $W(x, y)$  is increasing in  $x$  and increasing in  $y$  then,

$$\int_0^\infty \int_0^\infty W(x, y) \bar{F}(x, y) dx dy \leq \int_0^\infty \int_0^\infty W(x, y) \bar{G}(x, y) dx dy$$

**Proof.** Suppose  $W$  is increasing and  $F$  is not identically equal to  $G$ .

Since  $F$  is BIFR and  $G$  is bivariate exponential distribution with common mean  $\sim$ .

$\bar{F}$  crosses  $\bar{G}$  exactly once from the above,Say at  $(t_0, s_0)$ . That is ,

$$\bar{F}(t_0, s_0) = \bar{G}(t_0, s_0).$$

Now,

$$\begin{aligned} & \int_0^\infty \int_0^\infty w(x, y) \bar{F}(x, y) dx dy - \int_0^\infty \int_0^\infty w(x, y) \bar{G}(x, y) dx dy \\ & \leq \int_0^\infty \int_0^\infty [w(x, y) - w(t_0, s_0)] [\bar{F}(x, y) - \bar{G}(x, y)] dx dy \\ & \leq 0 \end{aligned}$$

This proves the theorem.

**Theorem 3.2** If  $F$  is BIFR with  $r$ -th moment  $\sim_{(r)}$ , then

$$\sim_{(r)} \begin{cases} \leq \sim^r r! & : r \geq 1, \\ \geq \sim^r r! & : 0 \leq r \leq 1 \end{cases}$$

**Proof.** Put  $w(x, y) = (xy)^{r-1}$  in the previous theorem then

$$\begin{aligned} & \int_0^\infty \int_0^\infty (xy)^{r-1} \bar{F}(x, y) dx dy \\ & \leq \int_0^\infty \int_0^\infty (xy)^{r-1} \bar{G}(x, y) dx dy \text{ for } r \geq 1 \\ (i.e.) \frac{\sim_{(r)}}{r} & \leq \int_0^\infty \int_0^\infty (xy)^{r-1} e^{-\frac{(x+y)}{\sim}} dx dy \\ & \leq \left[ \int_0^\infty (x)^{r-1} e^{-\frac{(x)}{\sim}} dx \right] \left[ \int_0^\infty (y)^{r-1} e^{-\frac{(y)}{\sim}} dy \right] \\ & \leq \sim^r (r-1)! \\ \sim_{(r)} & \leq r! \sim^r \end{aligned}$$

This completes the proof of the theorem.

**Theorem 3.3** If  $F$  is BNBUE then for all integers  $r \geq 0$

$$\frac{\sim(r+2)}{(r+2)} \leq \frac{\sim \cdot \sim(r+1)}{(r+1)}$$

**Proof.** Since  $F$  is BNBUE, we have

$$\begin{aligned} \bar{W}(x, y) &= \int_x^\infty \int_y^\infty \bar{F}(u, v) dv du. \text{ Thus for all integers } r > 0 \\ \int_0^\infty \int_0^\infty x^r y^r \bar{W}(x, y) dy dx &\leq \sim \int_0^\infty \int_0^\infty x^r y^r \bar{F}(x, y) dy dx \\ &\leq \sim \frac{\sim_{(r+1)}}{(r+1)^2} \\ &\leq \sim \frac{\sim_{(r+1)}}{(r+1)} \end{aligned} \tag{3.1}$$

Further

$$\begin{aligned} & \int_0^\infty \int_0^\infty x^r y^r \bar{W}(x, y) dy dx \\ &= E \left[ \int_0^\infty \int_0^\infty x^r y^r (X-x)(Y-y) I(X>x) I(Y>y) dy dx \right] \\ &= E \left[ \left( \int_0^\infty x^r (X-x) I(X>x) dx \right) \left( \int_0^\infty y^r (Y-y) I(Y>y) dy \right) \right] \\ &= E \left[ \left( X \int_0^X x^r dx \right) \left( Y \int_0^y y^r dy - \int_0^y y^{r+1} dy \right) \right] \end{aligned}$$

$$\begin{aligned} &= E \left[ \left( X \frac{x^{r+1}}{r+1} \Big|_0^X - \frac{x^{r+2}}{r+2} \Big|_0^X \right) \left( Y \frac{y^{r+1}}{r+1} \Big|_0^Y - \frac{y^{r+2}}{r+2} \Big|_0^Y \right) \right] \\ &= E \left[ X^{r+2} \left( \frac{1}{r+1} - \frac{1}{r+2} \right) \times Y^{r+2} \left( \frac{1}{r+1} - \frac{1}{r+2} \right) \right] \\ &= E \left[ X^{r+2} Y^{r+2} \left[ \frac{r+2-r-1}{(r+1)(r+2)} \right] \right] \\ &= \frac{\sim_{(r+2)}}{(r+1)(r+2)} \end{aligned}$$

Therefore the inequality (3.1) becomes

$$\begin{aligned} \frac{\sim_{(r+2)}}{(r+1)(r+2)} &\leq \frac{\sim \cdot \sim_{(r+1)}}{(r+1)^2} \\ \frac{\sim_{(r+2)}}{(r+2)} &\leq \frac{\sim \cdot \sim_{(r+1)}}{(r+1)} \end{aligned}$$

This completes the proof of the theorem.

**Theorem 3.4** If  $F$  is BHNBUUE then for all integers  $r \geq 0$

$$\sim_2 \leq 2 \sim^2$$

**Proof.** Since  $F$  is BHNBUUE we have

$$\bar{W}(x, y) \leq \sim \exp(-\frac{x+y}{\sim})$$

$$\begin{aligned} \int_0^\infty \int_0^\infty x^r y^r \bar{W}(x, y) dy dx &\leq \sim \int_0^\infty \int_0^\infty x^r y^r \exp(-\frac{x+y}{\sim}) dy dx \\ &= \sim^{2r+3} (r!)^2 \end{aligned}$$

But

$$\int_0^\infty \int_0^x x^r y^r \bar{W}(x, y) dy dx = \frac{\sim_{(r+2)}}{(r+1)(r+2)}$$

Therefore

$$\begin{aligned} \frac{\sim_{(r+2)}}{(r+1)(r+2)} &\leq \sim^{r+2} (r!) \\ \frac{\sim_{(r+2)}}{(r+2)!} &\leq \sim^{r+2} \end{aligned}$$

If  $r = 0$

$$\begin{aligned} \frac{\sim_2}{2!} &= \sim^2 \\ \sim_2 &= 2 \sim^2 \end{aligned}$$

This completes the proof of the theorem.

**Theorem 3.5** Let  $F$  be BNBUFR such that for some integers  $r, s \geq 0$

$$\frac{\sim_{r+s+2}}{(r+s+2)} \leq \frac{\sim_{(r+1)}}{(r+1)} \frac{s!}{[r(0,0)]^{s+1}}$$

**Proof.** Since  $F$  is BNBUFR, we have

$$\bar{F}(x+u, y+v) \leq \bar{F}(x, y) e^{-r(0,0)\sqrt{u^2+v^2}}.$$

Then for all integers  $r, s \geq 0$ ,

$$\int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty x^r y^r u^s v^s \bar{F}(x+u, y+v) dv du dy dx$$

$$\begin{aligned}
 &\leq \left( \int_0^\infty \int_0^\infty x^r y^r \bar{F}(x, y) dy dx \right) \\
 &\quad \left( \int_0^\infty \int_0^\infty u^s v^s e^{-r} (0, 0)^{\sqrt{u^2 + v^2}} dv du \right) \\
 &\leq E \left[ \int_0^X \int_0^Y x^r y^r dy dx \right] \frac{1}{r(0,0)^{s+1}} (s!) \\
 &\leq \frac{\tilde{(r+1)}}{r+1} \frac{s!}{[r(0,0)^{s+1}]} 
 \end{aligned}
 \tag{3.2}$$

Also

$$\begin{aligned}
 &\int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty x^r y^r u^s v^s \bar{F}(x+u, y+v) dv du dy dx \\
 &= \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (xy)^r (uv)^s \bar{F}(x+u, y+v) dv du dy dx \\
 &= \int_0^\infty \int_0^\infty \int_0^r \int_0^s (r-s)^r s^s (x-u)^r u^s \bar{F}(r, s) dx dr ds \\
 &= \int_0^\infty \int_0^\infty r^{r+s+1} s^{r+s+1} \bar{F}(r, s) dx dr \\
 &\quad \int_0^1 \int_0^1 (pq)^s (1-pq)^r dq dp \\
 &= s(r+1, s+1) \int_0^\infty \int_0^\infty (rx)^{r+s+1} \bar{F}(r, s) dx dr \\
 &= \frac{\tilde{(r+s+2)}}{(r+s+2)} \frac{r! s!}{(r+s+1)}
 \end{aligned}$$

Therefore the inequality (3.3) becomes

$$\frac{\tilde{(r+s+2)}}{(r+s+2)} \leq \frac{\tilde{(r+1)}}{(r+1)} \frac{(s!)}{[r(0,0)^{s+1}]}$$

This completes the proof of the theorem.

**Theorem 3.6** Suppose that  $F$  is BNBRU life distribution and it's  $(r+3)^{rd}$  moment is finite for some integer  $r \geq 0$ , then

$$\begin{aligned}
 &\frac{\tilde{(m+3)}}{(m+2)(m+3)} \\
 &\leq \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(i+1)} \tilde{(m-i+2)}}{(i+1)(m-i+1)(m-i+2)}
 \end{aligned}$$

**Proof.** Since  $F$  is a BNBRU, we have

$$\bar{W}(x, y) = \frac{1}{\sim} \int_y^\infty \int_x^\infty \bar{F}(u, v) dv du$$

Then the above inequality becomes

$$\bar{W}(x+t, y+s) \leq \bar{F}(x, y) \bar{W}(t, s) \forall x, y, t, s > 0$$

Multiplying both side by  $(x+t)^m (y+s)^m$  and integrating we get

$$\begin{aligned}
 &\int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{W}(x+t, y+s) dx dy ds dt \\
 &\leq \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x, y) \bar{W}(t, s) dx dy ds dt
 \end{aligned}
 \tag{3.3}$$

Also

$$\begin{aligned}
 &\int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{W}(x+t, y+s) dx dy ds dt \\
 &= \int_0^\infty \int_0^\infty \int_s^\infty \int_t^\infty (z_1 z_2)^m \bar{W}(z_1, z_2) dz_1 dz_2 ds dt
 \end{aligned}$$

$$\begin{aligned}
 &= \int_0^\infty \int_0^r (rs)^m \bar{W}(r, s) \int_0^s \int_0^r dz_1 dz_2 dr ds \\
 &= \int_0^\infty \int_0^\infty (rs)^{m+1} \bar{W}(r, s) dr ds \\
 &= \frac{1}{\sim} \int_0^\infty \int_0^\infty (rs)^{m+1} \int_s^\infty \int_r^\infty \bar{F}(u, v) dv du dr ds \\
 &= \frac{1}{\sim} \int_0^\infty \int_0^\infty \bar{F}(u, v) \int_u^\infty \int_0^v (rs)^{m+1} dr ds dv du \\
 &= \frac{1}{\sim} \int_0^\infty \int_0^\infty \frac{(uv)^{m+2}}{m+2} \bar{F}(u, v) dv du \\
 &= \frac{1}{\sim(m+2)} \int_0^\infty \int_0^\infty (uv)^{m+2} \bar{F}(u, v) dv du \\
 &= \frac{\tilde{(m+3)}}{\sim(m+2)(m+3)}
 \end{aligned}
 \tag{3.4}$$
  

$$\begin{aligned}
 &\text{Also } \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x, y) \bar{W}(t, s) dx dy ds dt \\
 &= \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x, y) \bar{W}(t, s) dx dy ds dt \\
 &= \sum_{i=0}^m \binom{m}{i} \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty x^i t^{m-i} y^i s^{m-i} \bar{F}(x, y) \bar{W}(t, s) dx dy ds dt \\
 &= \sum_{i=0}^m \binom{m}{i} \left[ \begin{array}{l} \left( \int_0^\infty \int_0^\infty (xy)^i \bar{F}(x, y) dx dy \right) \\ \left( \int_0^\infty \int_0^\infty (ts)^{m-i} \bar{W}(t, s) ds dt \right) \end{array} \right] \\
 &= \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(i+1)}}{(i+1)} \left[ \frac{1}{\sim} \int_0^\infty \int_0^\infty (ts)^{m-i} \int_s^\infty \int_t^\infty (uv)^{m-i} dv du ds dt \right] \\
 &= \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(i+1)}}{(i+1)} \frac{1}{\sim} \int_0^\infty \int_0^\infty \frac{(ts)^{m-i+1}}{(m-i+1)} \bar{F}(t, s) ds dt \\
 &= \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(i+1)}}{(i+1)} \frac{1}{\sim} \frac{\tilde{(m-i+2)}}{(m-i+1)(m-i+2)}
 \end{aligned}
 \tag{3.5}$$

Using (3.4) and (3.5) in (3.3) we get

$$\frac{\tilde{(m+3)}}{(m+2)(m+3)} = \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(i+1)} \tilde{(m-i+2)}}{(i+1)(m-i+1)(m-i+2)}$$

This completes the proof of the theorem.

**Theorem 3.7** For all non-negative integer  $r \geq 0$  and  $F$  is BNBRBU we get

$$\frac{\tilde{(m+2)}}{(m+2)} \leq \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(m-i+1)}}{(m-i+1)} \frac{\tilde{(i+2)}}{(i+1)(i+2)}$$

**Proof.** Since  $F$  is a BNBRBU, we have

$$\bar{W}(t, s) = \int_s^\infty \int_t^\infty \bar{F}(u, v) dv du$$

Multiplying both sides by  $[(x+t)(y+s)]^m$  and integrating, we have

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x+t, y+s) dx dy dt ds \\ & \leq \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x, y) \bar{W}(t, s) dx dy dt ds \end{aligned} \quad (3.6)$$

Also

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x+t, y+s) dx dy dt ds \\ & = \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x+t, y+s) dx dy dt ds \\ & = \int_0^\infty \int_0^\infty \int_0^X \int_0^Y (uv)^m \bar{F}(t, s) dv du dt ds \\ & = E \left[ \int_0^X \int_0^Y (xy)^{m+1} I[X - x, Y - y] dx dy \right] \\ & = \frac{E(XY)^{m+2}}{m+2} \\ & = \frac{\tilde{(m+2)}}{(m+2)} \end{aligned} \quad (3.7)$$

Also

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x, y) \bar{W}(t, s) dx dy dt ds \\ & = \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{F}(x, y) \bar{W}(t, s) dx dy dt ds \\ & = \sum_{i=0}^m \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty \binom{m}{i} x^{m-i} t^i y^{m-i} s^i \bar{F}(x, y). \end{aligned}$$

$$\begin{aligned} & \left( \int_s^\infty \int_t^\infty \bar{F}(u, v) dv du \right) dx dy dt ds \\ & = \sum_{i=0}^m \binom{m}{i} \int_0^\infty \int_0^\infty (xy)^{m-1} \bar{F}(x, y) dx dy. \\ & \int_0^\infty \int_0^\infty (ts)^i \int_s^\infty \int_t^\infty \bar{F}(u, v) dv du ds dt \\ & = \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(m-i+1)}}{(m-i+1)} \int_0^\infty \int_0^\infty (ts)^i \int_s^\infty \int_t^\infty \bar{F}(u, v) dv du ds dt \\ & = \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(m-i+1)}}{(m-i+1)} \int_0^\infty \int_0^\infty \bar{F}(x, y) \int_0^X \int_0^Y (u, v)^i dv du dx dy \\ & = \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(m-i+1)}}{(m-i+1)} \int_0^\infty \int_0^\infty \frac{(XY)^{i+1}}{(i+1)} \bar{F}(x, y) dx dy \\ & = \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(m-i+1)}}{(m-i+1)} \frac{1}{(i+1)} \frac{\tilde{(i+2)}}{(i+2)} \\ & = \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(m-i+1)}}{(m-i+1)} \frac{\tilde{(i+2)}}{(i+1)(i+2)} \end{aligned} \quad (3.8)$$

Using (3.7) and (3.8) the inequality (3.6) becomes

$$\frac{\tilde{(m+2)}}{(m+2)} \leq \sum_{i=0}^m \binom{m}{i} \frac{\tilde{(m-i+1)}}{(m-i+1)} \frac{\tilde{(i+2)}}{(i+1)(i+2)}$$

This completes the proof of the theorem.

**Theorem 3.8** For all non-negative integer  $r \geq 0$  and  $F$  is BRNU we get

$$\frac{\tilde{(m+3)}}{(m+2)(m+3)} \leq \frac{1}{\sim} \sum_{i=0}^m \binom{m}{i} \frac{1}{(m-i+1)(i+1)} \frac{\tilde{(m-i+2)} \tilde{(i+2)}}{(m-i+2)(i+2)}$$

**Proof.** Since  $F$  is a BRNU, we have

$$\bar{W}(x+t, y+s) \leq \bar{W}(x, y) \bar{W}(t, s)$$

Multiplying both sides by  $(x+t)^m (y+s)^m$  and integrating over  $(0, \infty)$ , we have

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{W}(x+t, y+s) dx dy ds dt \\ & \leq \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{W}(x, y) \bar{W}(t, s) dx dy ds dt \\ & \text{Also} \\ & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{W}(x+t, y+s) dx dy ds dt \\ & = \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{W}(x+t, y+s) dx dy ds dt \\ & = \frac{1}{\sim} \int_0^\infty \int_0^\infty \bar{F}(x, y) \int_0^X \int_0^Y (uv)^{m+1} dv du dx dy \\ & = \frac{1}{\sim} E \left[ \int_0^\infty \int_0^\infty \frac{(XY)^{m+2}}{m+2} I[X > x, Y > y] dx dy \right] \\ & = \frac{1}{\sim} \frac{1}{m+2} E(xy)^{m+2} \\ & = \frac{\tilde{(m+3)}}{\sim(m+2)(m+3)} \end{aligned} \quad (3.10)$$

where  $I$  is a indicator function.

Also

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{W}(x, y) \bar{W}(t, s) dx dy ds dt \\ & = \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (x+t)^m (y+s)^m \bar{W}(x, y) \bar{W}(t, s) dx dy ds dt \\ & = \frac{1}{\sim^2} \sum_{i=0}^m \binom{m}{i} \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty x^{m-i} t^i y^{m-i} s^i \\ & \quad \bar{W}(x, y) \bar{W}(t, s) dx dy ds dt \\ & = \frac{1}{\sim^2} \sum_{i=0}^m \binom{m}{i} \left( \int_0^\infty \int_0^\infty (xy)^{m-i} \bar{W}(x, y) dx dy \right) \\ & \quad \left( \int_0^\infty \int_0^\infty (ts)^i \bar{W}(t, s) ds dt \right) \\ & = \frac{1}{\sim^2} \sum_{i=0}^m \binom{m}{i} \int_0^\infty \int_0^\infty \frac{(xy)^{m-i+1}}{(m-i+1)} I(X > x, Y > y) dx dy \\ & \quad E \left[ \int_0^\infty \int_0^\infty \frac{(ts)^{i+1}}{i+1} I(T > t, S > s) ds dt \right] \\ & = \frac{1}{\sim^2} \sum_{i=0}^m \binom{m}{i} \frac{1}{(m-i+1)(i+1)} \frac{\tilde{(m-i+2)}}{(m-i+2)} \frac{\tilde{(i+2)}}{(i+2)} \end{aligned} \quad (3.11)$$

Using (3.10) and (3.11) the inequality (3.9) become

$$\begin{aligned} & \frac{\tilde{(m+3)}}{(m+2)(m+3)} = \\ & \frac{1}{\sim} \sum_{i=0}^m \binom{m}{i} \frac{1}{(m-i+1)(i+1)} \frac{\tilde{(m-i+2)}}{(m-i+2)} \frac{\tilde{(i+2)}}{(i+2)} \end{aligned}$$

This completes the proof of the theorem.

**Theorem 3.9** If  $F$  is BRNBRU then

$$\frac{\sim \cdot \sim_{(m+n+3)}}{(m+n+3)!} \leq \frac{\sim_{(m+2)} \sim_{(n+2)}}{(m+2)!(n+2)!}$$

**Proof.** Let

$$\bar{V}(x, y) = \int_x^\infty \int_y^\infty \bar{F}(u, v) dv du.$$

Since  $F$  is a BRNBRU, we have

$$\sim \bar{V}(x+t, y+s) \leq \bar{V}(x, y) \bar{V}(t, s)$$

Multiplying by  $(xy)^m (ts)^n ; m, n > 0$  (integers) and integrating over  $(0, \infty)$  with respect to  $x, y, t, s$  we get,

$$\begin{aligned} & \sim \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (xy)^m (ts)^n \bar{V}(x+t, y+s) dx dy ds dt \\ & \leq \left( \int_0^\infty \int_0^\infty (xy)^m \bar{V}(x, y) dx dy \right) \left( \int_0^\infty \int_0^\infty (ts)^n \bar{V}(t, s) ds dt \right) \end{aligned} \quad (3.12)$$

Also

$$\begin{aligned} & \sim \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (xy)^m (ts)^n \bar{V}(x+t, y+s) dx dy ds dt \\ & = \sim \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (xy)^m (ts)^n \bar{V}(x+t, y+s) dx dy ds dt \\ & = \frac{\sim}{(m+1)(m+2)} E \left[ \int_0^\infty \int_0^\infty (ts)^n E(XY - ts)^{m+2} I(X > t, Y > s) ds dt \right] \\ & = \frac{\sim}{(m+1)(m+2)} E \left[ \int_0^X \int_0^Y (ts)^n (XY - ts)^{m+2} ds dt \right] \\ & = \frac{\sim}{(m+1)(m+2)} E \left[ (XY)^{m+n+3} \int_0^1 \int_0^1 (rs)^n (1-rs)^{m+2} dr ds \right] \\ & = \frac{\sim}{(m+1)(m+2)} E \left[ (XY)^{m+n+3} \right] \frac{\Gamma(x+1)\Gamma(m+3)}{\Gamma(m+n+4)} \\ & = \frac{\sim \cdot \sim_{(m+n+3)}}{(m+1)(m+2)} \frac{n!(m+2)!}{(m+n+3)!} \\ & = \frac{m!n! \sim \cdot \sim_{(m+n+3)}}{(m+n+3)!} \end{aligned} \quad (3.13)$$

Consider

$$\begin{aligned} & \int_0^\infty \int_0^\infty (xy)^m \bar{V}(x, y) dx dy = \\ & \quad E \left[ \int_0^\infty \int_0^\infty (xy)^m (XY - xy) I(X > x, Y > y) dy dx \right] \\ & = E \left[ XY \int_0^X \int_0^Y (xy)^m dy dx - \int_0^X \int_0^Y (xy)^{m+1} dy dx \right] \\ & = E \left[ \frac{(XY)^{m+2}}{(m+1)} - \frac{(XY)^{m+2}}{(m+2)} \right] \\ & = E \left[ (XY)^{m+2} \left( \frac{1}{(m+1)} - \frac{1}{(m+2)} \right) \right] \\ & = \frac{1}{(m+1)(m+2)} E[(XY)^{m+2}] \end{aligned}$$

$$\int_0^\infty \int_0^\infty (xy)^m \bar{V}(x, y) dy dx = \frac{\sim_{(m+2)}}{(m+1)(m+2)} \quad (3.14)$$

Similarly

$$\int_0^\infty \int_0^\infty (ts)^n \bar{V}(t, s) ds dt = \frac{\sim(n+2)}{(n+1)(n+2)} \quad (3.15)$$

Using (3.13), (3.14) and (3.15) in (3.12) we get

$$\begin{aligned} & \frac{m!n! \sim \cdot \sim_{(m+n+3)}}{(m+n+3)!} \leq \frac{\sim_{(m+2)}}{(m+1)(m+2)} \frac{\sim_{(n+2)}}{(n+1)(n+2)} \\ & \frac{\sim \cdot \sim_{(m+n+3)}}{(m+n+3)!} \leq \frac{\sim_{(m+2)} \sim_{(n+2)}}{(m+2)!(n+2)!} \end{aligned}$$

This completes the proof of the theorem.

**Theorem 3.10** If  $F$  is BRNBRUE then

$$2 \sim \frac{\sim_{(m+2)}}{(m+1)!} \leq \frac{\sim_{(2)} \sim_{(m+3)}}{(m+3)}$$

**Proof.** If  $F$  is a BRNBRUE, we have

$$2 \sim \int_y^\infty \int_x^\infty \bar{W}(u, v) dv du \leq \sim_{(2)} \bar{W}(x, y)$$

Multiplying both sides by  $(xy)^r$  and inequality

$$\begin{aligned} & 2 \sim \int_0^\infty \int_0^\infty \int_y^\infty \int_x^\infty (xy)^m \bar{W}(u, v) dv du dx dy \\ & \leq \sim_{(2)} \int_0^\infty \int_0^\infty (xy)^m \bar{W}(x, y) dx dy \end{aligned} \quad (3.16)$$

Consider

$$\begin{aligned} & \int_0^\infty \int_0^\infty (xy)^m \bar{W}(x, y) dx dy = \\ & \quad \int_0^\infty \int_0^\infty (xy)^m \left( \int_y^\infty \int_x^\infty \bar{F}(u, v) du dv \right) dx dy \\ & = \int_0^\infty \int_0^\infty \bar{F}(x, y) \int_0^y \int_0^x (uv)^m du dv dx dy \\ & = \int_0^\infty \int_0^\infty \frac{(xy)^{m+1}}{(m+1)^2} \bar{F}(x, y) dx dy \\ & = \frac{1}{(m+1)^2} \int_0^\infty \int_0^\infty (xy)^{m+1} \bar{F}(x, y) dx dy \end{aligned}$$

$$\int_0^\infty \int_0^\infty (xy)^m \bar{W}(x, y) dx dy = \frac{1}{(m+1)^2} \frac{\sim_{(m+2)}}{(m+2)} \quad (3.17)$$

and

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_y^\infty \int_x^\infty (xy)^m \bar{W}(u, v) dv du dx dy \\ & = \int_0^\infty \int_0^\infty \bar{W}(x, y) \frac{(xy)^{m+1}}{m+1} dx dy \\ & = \frac{1}{(m+1)} E \left[ \int_0^\infty \int_0^\infty (xy)^{m+1} (XY - xy) I(X > x, Y > y) dx dy \right] \\ & = \frac{1}{(m+1)} E \left[ XY \int_0^Y \int_0^X (xy)^{m+1} dy dx - \int_0^Y \int_0^X (xy)^{m+2} dy dx \right] \end{aligned}$$

$$\begin{aligned}
 &= \frac{1}{(m+1)} E \left[ \frac{(XY)^{m+2}}{m+2} - \frac{(XY)^{m+3}}{m+3} \right] \\
 &= \frac{1}{(m+1)} \tilde{\}_{(m+3)} \left( \frac{1}{(m+2)} - \frac{1}{(m+3)} \right) \\
 &= \frac{\tilde{\}_{(m+3)}}{(m+1)(m+2)(m+3)}
 \end{aligned} \tag{3.18}$$

Using (3.17) and (3.18) the inequality (3.16) becomes

$$\begin{aligned}
 2 \sim \frac{\tilde{\}_{(m+2)}}{(m+1)^2(m+2)} &\leq \frac{\tilde{\}_{(2)} \tilde{\}_{(m+3)}}{(m+1)(m+2)(m+3)} \\
 2 \sim \frac{\tilde{\}_{(m+2)}}{(m+1)} &\leq \frac{\tilde{\}_{(2)} \tilde{\}_{(m+3)}}{(m+3)}
 \end{aligned}$$

This completes the proof of the theorem.

**Theorem 3.11** Let  $W(\{\}) = \int_0^\infty \int_0^\infty e^{-\{\}(x+y)} dF(x+y)$

if  $F$  is BNBUL. Then for all integer  $s, r \geq 0$

$$\begin{aligned}
 &\frac{(-1)^{m+1} m!}{\{\}^{m+1}} [1 - \{\}] + \\
 &\frac{m!}{\{\}^{m+1}} \sum_{i=0}^m \frac{(-1)^i}{(m-i)!} \{\}^{m-i} \frac{\tilde{\}_{(m-i+1)}}{(m-i+1)} \\
 &= \frac{\tilde{\}_{(m+1)}}{(m+1)} \frac{(1 - \{\})}{\{\}}.
 \end{aligned}$$

**Proof.** Let

$$W(\{\}) = \int_0^\infty \int_0^\infty e^{-\{\}(x+y)} dF(x, y).$$

Since  $F$  is BNBUL we have

$$\begin{aligned}
 &\int_0^\infty \int_0^\infty e^{-\{\}(x+y)} \bar{F}(x+t, y+s) dx dy \\
 &\leq \bar{F}(t, s) \int_0^\infty \int_0^\infty e^{-\{\}(x+y)} \bar{F}(x, y) dy dx
 \end{aligned}$$

Multiplying both side by  $(ts)^m$  and integrating over  $(0, \infty)$ , we obtain,

$$\begin{aligned}
 &\int_0^\infty \int_0^\infty (ts)^m \int_0^\infty \int_0^\infty e^{-\{\}(x+y)} \bar{F}(x+t, y+s) dy dx ds dt \\
 &\leq \int_0^\infty \int_0^\infty (ts)^m \bar{F}(t, s) ds dt \int_0^\infty \int_0^\infty e^{-\{\}(x+y)} \bar{F}(x, y) dy dx
 \end{aligned} \tag{3.19}$$

Consider

$$\begin{aligned}
 \int_0^\infty \int_0^\infty (ts)^m \bar{F}(t, s) ds dt &= E \left[ \int_0^\infty \int_0^\infty (ts)^m I(T > t, S > s) ds dt \right] \\
 &= \frac{\tilde{\}_{(m+1)}}{(m+1)}
 \end{aligned} \tag{3.20}$$

$$\int_0^\infty \int_0^\infty e^{-\{\}(x+y)} \bar{F}(x, y) dy dx = \int_0^\infty \int_0^\infty e^{-\{\}(x+y)} (1 - F(x, y)) dy dx$$

$$= \frac{1}{\{\}} - \int_0^\infty \int_0^\infty e^{-\{\}(x+y)} F(x, y) dy dx$$

$$= \frac{1}{\{\}} - \frac{1}{\{\}} W(\{\})$$

$$\begin{aligned}
 &= \frac{1}{\{\}} (1 - W(\{\})) \\
 &= \frac{\tilde{\}_{(m+1)}}{m+1} \frac{1}{\{\}} [1 - W(\{\))]
 \end{aligned} \tag{3.21}$$

Using (3.20) and (3.21) in the equation (3.19) we have

$$= \frac{\tilde{\}_{(m+1)}}{m+1} \frac{1}{\{\}} [1 - W(\{\))] \tag{3.22}$$

$$\begin{aligned}
 &= \int_0^\infty \int_0^\infty (ts)^m \int_0^\infty \int_0^\infty e^{-\{\}(x+y)} \bar{F}(x+t, y+s) dy dx ds dt \\
 &= \int_0^\infty \int_0^\infty e^{-\{\}(t+s)} \bar{F}(t, s) \int_0^s \int_0^t (u+v)^m e^{\{\}(u+v)} dv du ds dt
 \end{aligned} \tag{3.23}$$

Consider

$$\begin{aligned}
 &\int_0^s \int_0^t (u+v)^m e^{\{\}(u+v)} dv du = \\
 &\frac{m}{\{\}^{m+1}} \left[ (-1)^{m+1} + \sum_{i=0}^m (-1)^i \frac{[\{\} (s+t)]^{m-i}}{(m-i)!} e^{\{\}(s+t)} \right]
 \end{aligned} \tag{3.24}$$

Therefore (3.23) becomes

$$\begin{aligned}
 &= \int_0^\infty \int_0^\infty e^{-\{\}(t+s)} \bar{F}(t, s) \frac{m!}{\{\}^{m+1}} \\
 &\quad \left[ (-1)^{m+1} + \sum_{i=0}^m (-1)^i \frac{[\{\} (t+s)]^{m-i}}{(m-i)!} e^{\{\}(s+t)} \right] ds dt \\
 &= \frac{(-1)^{m+1} m!}{\{\}^{m+1}} \left[ \int_0^\infty \int_0^\infty e^{-\{\}(t+s)} \bar{F}(t, s) ds dt \right] \\
 &\quad + \frac{m!}{\{\}^{m+1}} \sum_{i=0}^m \frac{(-1)^i}{(m-i)!} \int_0^\infty \int_0^\infty [\{\}^{m-i} (t+s)^{m-i}] \bar{F}(t, s) ds dt \\
 &= \frac{(-1)^{m+1} m!}{\{\}^{m+1}} [1 - \{\}] + \frac{m!}{\{\}^{m+1}} \sum_{i=0}^m \frac{(-1)^i}{(m-i)!} \\
 &\quad \{\}^{m-i} \frac{\tilde{\}_{(m-i+1)}}{(m-i+1)}
 \end{aligned} \tag{3.25}$$

Using (3.22) and (3.25) then the inequality (3.19) becomes

$$\begin{aligned}
 &\frac{(-1)^{m+1} m!}{\{\}^{m+1}} [1 - \{\}] + \\
 &\frac{m!}{\{\}^{m+1}} \sum_{i=0}^m \frac{(-1)^i}{(m-i)!} \{\}^{m-i} \frac{\tilde{\}_{(m-i+1)}}{(m-i+1)} \\
 &= \frac{\tilde{\}_{(m+1)}}{(m+1)} \frac{(1 - \{\})}{\{\}}
 \end{aligned}$$

This completes the proof of the theorem.

**Theorem 3.12** Let  $F$  be a life distribution which is BEBU with mean  $\sim$  then

$$\tilde{\}_{n+m}} \leq \{\}_{m+n}$$

**Proof.** We shall consider only the BEBU case

$$\bar{F}(x+t, y+s) \leq \bar{F}(t, s) e^{\frac{(x+y)}{\{\}}}$$

Multiplying both sides by  $\frac{(xy)^{m-1} (ts)^{n-1}}{\Gamma(m)\Gamma(n)}$  and integrating

over  $(0, \infty)$ , we obtain,

$$\int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty \frac{(xy)^{m-1} (ts)^{n-1}}{\Gamma(m)\Gamma(n)} \bar{F}(x+t, y+s) dx dy ds dt$$

$$= \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty \frac{(xy)^{m-1} (ts)^{n-1}}{\Gamma(m)\Gamma(n)} \overline{F}(t,s) e^{-\frac{(x+y)}{\sim}} dx dy ds dt$$

(3.26)

$$= \left( \int_0^\infty \int_0^\infty \frac{(xy)^{m-1}}{\Gamma(m)} e^{-\frac{(x+y)}{\sim}} dx dy \right) \left( \int_0^\infty \int_0^\infty \frac{(ts)^{n-1}}{\Gamma(n)} \overline{F}(t,s) ds dt \right)$$

$$= \sim^m \} _n \quad (3.27)$$

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty \frac{(xy)^{m-1} (ts)^{n-1}}{\Gamma(m)\Gamma(n)} \overline{F}(x+t, y+s) dx dy ds dt \\ & = \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty \frac{(xy)^{m-1} (ts)^{n-1}}{\Gamma(m)\Gamma(n)} \int_{x+t}^\infty \int_{y+s}^\infty dF(u,v) dx dy ds dt \end{aligned}$$

On changing the variables we get,

$$\begin{aligned} & = \int_0^\infty \int_0^\infty \left( \int_0^{z_1} \int_0^{z_2} \frac{(xy)^{m-1}}{\Gamma(m)} \right. \\ & \quad \left. \left( \int_0^{z_1-x} \int_0^{z_2-y} \frac{(ts)^{n-1}}{\Gamma(n)} ds dt \right) dF(z_1, z_2) \right) \\ & = \int_0^\infty \int_0^\infty \int_0^{z_1} \int_0^{z_2} \frac{(xy)^{m-1}}{\Gamma(m)} \frac{(z_1-x)^n (z_2-y)^n}{\Gamma(n+1)} \\ & \quad dx dy dF(z_1, z_2) \\ & = \int_0^\infty \int_0^\infty \int_0^1 \int_0^1 \frac{(\Gamma z_1 \cdot S z_2)^{m-1} z_1^n (1-\Gamma)^n z_2^n (1-S)^n}{\Gamma(m)\Gamma(n+1)} \\ & \quad (z_1 d\Gamma)(z_2 dS) dF(z_1, z_2) \\ & = \int_0^\infty \int_0^\infty \int_0^1 \int_0^1 \frac{(z_1 z_2)^{m-1} z_1^{n+1} z_2^{n+1}}{\Gamma(m)\Gamma(n+1)} (\Gamma S)^{m-1} \\ & \quad [(1-\Gamma)(1-S)]^n d\Gamma dS dF(z_1, z_2) \\ & = \int_0^\infty \int_0^\infty \frac{z_1^{m+n} z_2^{m+n}}{\Gamma(m)\Gamma(n+1)} \left( \int_0^1 \int_0^1 (\Gamma S)^{m-1} [(1-\Gamma)(1-S)]^n d\Gamma dS \right) \\ & \quad dF(z_1, z_2) \\ & = \int_0^\infty \int_0^\infty \frac{(z_1 z_2)^{m+n}}{\Gamma(m)\Gamma(n+1)} \frac{\Gamma(m)\Gamma(n+1)}{\Gamma(m+n+1)} dF(z_1, z_2) \\ & = \frac{1}{\Gamma(m+n+1)} \int_0^\infty \int_0^\infty (z_1 z_2)^{m+n} dF(z_1, z_2) \\ & = \frac{E[(z_1 z_2)^{m+n}]}{\Gamma(m+n+1)} = \frac{\sim^{m+n}}{\Gamma(m+n+1)} \\ & = \} _{m+n}. \quad (3.28) \end{aligned}$$

Using (3.27) and (3.28) in (3.26) we get,

$$\sim^m \} _n \leq \} _{m+n}$$

This completes the proof of the theorem.

**Theorem 3.13** Suppose that  $F$  is BEBUC(2) life distribution such that its  $\sim_{r+s+4}$  the moment of order is finite ( $r+s+4$ ) for some integers  $r$  and  $s$  then the following moment inequality holds

$$\frac{\sim_{m+n+3}}{(m+n+3)!} \leq \frac{\sim_{(m+1)} \sim_{(n+2)}}{(n+2)!}$$

**Proof.** Since  $F$  is BEBUC

$$\int_u^\infty \int_v^\infty \overline{F}(x+t, y+s) ds dt \leq e^{-\frac{(x+y)}{\sim}} \int_u^\infty \int_v^\infty \overline{F}(t,s) ds dt \quad (3.29)$$

Multiplying both sides of (3.29) by  $(xy)^m (ts)^n$  and integrating we get

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty \int_u^\infty \int_v^\infty (xy)^m (ts)^n \overline{F}(x+t, y+s) dx dy ds dt dv du \\ & \leq \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty e^{-\left(\frac{x+y}{\sim}\right)} (xy)^m (ts)^n \int_u^\infty \int_v^\infty \overline{F}(t,s) ds dt dv du dx dy \end{aligned} \quad (3.30)$$

Also

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty e^{-\left(\frac{x+y}{\sim}\right)} (xy)^m (ts)^n \int_u^\infty \int_v^\infty \overline{F}(t,s) ds dt dv du dx dy \\ & = \int_0^\infty \int_0^\infty (xy)^m e^{-\left(\frac{x+y}{\sim}\right)} \int_0^\infty \int_0^\infty (ts)^n \int_u^\infty \int_v^\infty \overline{F}(t,s) ds dt dv du dy dx \\ & = m! \sim^{m+1} \int_0^\infty \int_0^\infty \overline{F}(ts) \int_0^u \int_0^v (t,s)^n dt ds du dv \\ & = \frac{m! \sim^{m+1}}{(n+1)} \int_0^\infty \int_0^\infty (uv)^{n+1} \overline{F}(t,s) ds dt \\ & = \frac{m! \sim^{m+1} \sim^{n+2}}{(n+1)(n+2)} \end{aligned} \quad (3.31)$$

Also

$$\begin{aligned} & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty \int_u^\infty \int_v^\infty (xy)^m (ts)^n \overline{F}(x+t, y+s) \\ & \quad dx dy ds dt dv du \\ & = \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty \int_u^\infty \int_v^\infty (xy)^m (ts)^n \overline{F}(x+t, y+s) \\ & \quad dx dy ds dt dv du \\ & = \frac{m! n! \sim_{(m+n+3)}}{(m+n+3)!} \end{aligned} \quad (3.32)$$

Using (3.31) and (3.32) in (3.30) we have

$$\frac{m! n! \sim_{(m+n+3)}}{(m+n+3)!} \leq \frac{m! \sim_{(m+1)} \sim_{(n+2)}}{(n+1)(n+2)}$$

$$\frac{\sim_{(m+n+3)}}{(m+n+3)!} \leq \frac{\sim_{(m+1)} \sim_{(n+2)}}{(n+2)!}$$

This completes the proof of the theorem.

**Theorem 3.14** If  $F$  is BEBUCA then

$$\frac{\sim_{(m+3)}}{(m+1)(m+2)(m+3)} \leq \sim^2 \sim_{(m+1)}$$

**Proof.** Since  $F$  is BEBUCA we have

$$\int_0^\infty \int_0^\infty \overline{W}(x+t, y+s) dx dy \leq \sim^2 \overline{F}(t,s).$$

Multiplying both sides by  $(ts)^m$  and integrating we get

$$\begin{aligned}
 & \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (ts)^m \bar{W}(x+t, y+s) dx dy dt ds \\
 & \leq \sim^2 \int_0^\infty \int_0^\infty (ts)^m \bar{F}(t, s) ds dt \\
 & = \int_0^\infty \int_0^\infty \int_0^\infty \int_0^\infty (ts)^m \bar{W}(x+t, y+s) dx dy dt ds \\
 & = \int_0^\infty \int_0^\infty \int_0^{z_2} \int_0^{z_1} \bar{W}(z_1 z_2) (ts)^m dt ds dz_1 dz_2 \\
 & = \int_0^\infty \int_0^\infty \bar{W}(z_1 z_2) \frac{t^{m+1}}{m+1} \Big|_0^{z_1} \frac{s^{m+1}}{m+1} \Big|_0^{z_2} dz_1 dz_2 \\
 & = \int_0^\infty \int_0^\infty \bar{W}(z_1 z_2) \frac{(z_1 z_2)^{m+1}}{(m+1)^2} dz_1 dz_2 \\
 & = \frac{1}{(m+1)^2} \int_0^\infty \int_0^\infty (z_1 z_2)^{m+1} \bar{W}(z_1 z_2) dz_1 dz_2 \\
 & = \frac{1}{(m+1)^2} \int_0^\infty \int_0^\infty (xy)^{m+1} \bar{W}(xy) dx dy \\
 & = \frac{1}{(m+1)^2} E \left[ \int_0^\infty \int_0^\infty (xy)^{m+1} (XY - xy) I(X > x, Y > y) dy dx \right] \\
 & \frac{\sim_{(m+3)}}{(m+1)(m+2)(m+3)} \leq \sim^2 \sim_{(m+1)}
 \end{aligned} \tag{3.33}$$

This completes the proof of the theorem.

## Conclusion

In this paper, we have derived the moment inequalities for bivariate ageing classes of life distributions.

## References

- [1] Abdel-Aziz, A. A., (2007), On testing exponentiality against RNBRUE Alternatives, *Applied Mathematical Sciences*, 1(35), 1725-1736.
  - [2] Abu-Youssef, S. E., (2015), On Properties of UBAC of life distribution, *Journal of Advances in Mathematics*, 10(3).
  - [3] Abu-Youssef, S. E., Elbatal, I. I., (2006), A note on moment inequality for hormonic used better than aged in expectation (HUBAE) class of life distribution with hypothesis testing application, *International Journal Comtemp. Maths. Sci.*, 1(3), 141-150.
  - [4] Ahmed, I. A., (2001), Moments inequality of Ageing Families of distributions with Hypothesis Testing Application, *Journal of Statistical Planning and Inference*, 92(12), 121-132.
  - [5] Ahmed, I., and Mugdadi, A., (2002), Further Moments inequality of Life distributions with Hypothesis Testing Application, *Journal of Statistical Planning and Inference*, 93, 121-132.
  - [6] Al-Ghafily, (2015), Moment inequality for exponential better than used in convex average class of life distribution with Hypothesis testing application, *Life Science Journal* 12(2).
  - [7] Al-Ruzaiza Hendi, M. I., and Abu-Youssef, (2003), A note on Moment inequality for the Harmonic New
  - =  $\frac{1}{(m+1)^2} E \left[ XY \int_0^X \int_0^Y (xy)^{m+1} dy dx - \int_0^X \int_0^Y (xy)^{m+2} dy dx \right]$
  - =  $\frac{1}{(m+1)^2} E \left[ \frac{(XY)^{m+3}}{(m+1)^2} \left( \frac{1}{m+2} - \frac{1}{m+3} \right) \right]$
  - =  $\frac{\sim_{(m+3)}}{(m+1)^2(m+2)(m+3)}$  (3.34)
  - =  $\sim^2 \int_0^\infty \int_0^\infty (ts)^m \bar{F}(t, s) ds dt$
  - =  $\sim^2 \frac{\sim_{(m+1)}}{(m+1)}$  (3.35)
- Using (3.34) and (3.35) in (3.33) we get
- $$\frac{\sim_{(m+3)}}{(m+1)^2(m+2)(m+3)} \leq \sim^2 \frac{\sim_{(m+1)}}{(m+1)}$$

Better than Used is expectation property with Hypothesis testing application, *Journal of Nonparametric Statistics* 15(3), 267-272.

- [8] Balakrishnan, N., and Chin-Diew Lai, (2009), *Continuous Bivariate Distribution*, Second Edition, Springer, New York.
- [9] Diah, L. S., El-Atfy, E.S., (2017), A moment inequality for over all decreasing life distribution with hypothesis testing application, *International Journal of Mathematics and Statics Invention (IJMSI)*, 62-71.
- [10] Diah, L. S., Kayid, M., and Aw.Mahamoud, M., (2009), Moments Inequality for NBUL Distribution with hypothesis testing application, *Contemporary Engineering Sciences*, 2(7), 319-332.
- [11] Elbatal, I., (2009), On testing Statistics of renewal new better than renewal used class of life distribution, *Int. Journal Contemp. Math. sciences*, 4(1), 17-29.
- [12] Ibrahim, B., Abdul-Moniem, (2011), Testing EBELC class of life distribution based on moments inequalities, *International Mathematics Forum*, 6(58), 2867-2879.
- [13] Ibrahim, B., Abdul-Moniem, (2008), Testing Exponentiality versus NBUASI based on moments inequalities, *Madina Higher Institute for Management and Technology*, Madina Academy, Giza, Egypt.
- [14] Ibrahim, B., Ahmed, A., (2001), Moments inequalities of aging families of distribution with hypothesis testing application, *Journal of Statistical Planning and Inference*, 92, 121-132.
- [15] Mahmod, M.A.W., Albassam, M.S., and Abdul fathah, E.H., (2010), A new approach to moment inequalities for NBRU class of life distributions with hypothesis testing applications, *International Journal of Reliability and Application*, 11(2), 139-151.
- [16] Mahmod, M.A.W., and Diab, S., (2013), A new approach for moments inequalities for NRBU and

- RNBU classes with hypothesis testing applications, *International Journal of Basic & Applied Sciences*, 13(6), 139-151.
- [17] Mahmud, M.A.W., El-Arishi, S. M., and Diab, L. S.,(2003), Moment Inequality for testing New Renewal Better than Used and Renewal New Better than Used classes, *International Journal of Reliability and Application*, 4, 97-123.
- [18] Nofal, Z.W., Hamed, M.S., (2014), On some new families based on the Exponential better than used concept, *Applied Mathematical Science*, 8(164), 8153-8170.
- [19] Barlow, Frank Proschan, and Larry C. Hunter,
- (1965), *Mathematical theory of reliability*, 33-43.
- [20] Rizwan U., (2000), On Stochastic Life time models, *Journal of Madras University Section B Sciences*, 52, 121-143.
- [21] Shokry (Cario) E. M., Ahmed (Cario) A. N., Rakha (Suez), and Hewedi (Suez),(2009), New result on the NBUFR and NBUE classes of life distribution, *Applied Mathematics*, 36(2), 139-147.
- [22] Syed Tahir Hussainy, (2015), on Some Ageing Properties of Bivariate Life distribution under convolution, *International Journal of Science and Humanities*, 1(1), 47-52.

\*\*\*\*\*